# EIGENVECTORS OF ELEMENTS IN $\mathbb{Z}[S U(2)]$ 

DANIEL CIZMA, JONATHAN GODIN, DMITRY JAKOBSON, AND LUIS LEDESMA VEGA

## 1. Introduction

In the paper GJS], the authors looked at eigenvalue distribution and level spacings distribution for the $N$-th irreducible representation of elements of the group ring of $S U(2)$ as $N \rightarrow \infty$.

In the current paper, we will study the behaviour of the corresponding eigenvectors.

## 2. Questions

Fix $k \geq 3$. Let $g=g_{1}+g_{1}^{-1}+\ldots+g_{k}+g_{k}^{-1}$ be a generic element of $\mathbb{R}[\operatorname{SU}(2)]$. Let $\pi_{N}$ denote the $N$-th irreducible representation of $\mathrm{SU}(2)$. It has been found in GJS that for $N$ odd, the eigenvalues of $\pi_{N}(g)$ are double for generic choice of $g \in \mathrm{SU}(2)^{k}$, while for $N$ even, the eigenvalues are simple for such $g$. We normalize eigenvectors to have $l^{2}$ norm equal to 1 .

For $N$ even, we get $(N+1)$ points on the unit sphere $S^{N} \subset \mathbb{R}^{N+1}$. It is a natural to ask the following:

Question 1. Do those points become uniformly distributed on $S^{N}$ as $N \rightarrow$ $\infty$ ?

A weaker question is whether this happens after averaging over $\mathrm{SU}(2)^{k}$.
For $N$ odd, the eigenvalues are double. Accordingly, a natural question seems to be the following:

Question 2. Are the 2-dimensional subspaces corresponding to the double eigenvalues becoming uniformly distributed in $\operatorname{Gr}(2, N+1)$, as $N \rightarrow \infty$ ?

## 3. Results for random matrices

Our motivation comes from the theory of random matrices, where the behaviour of eigenvectors was studied by Gaudin, Mehta, Knowles, Yin, Tao and Vu. We refer to [KnY, TV] and references therein for results about both Wigner matrices and general random matrices. One of the important results is that the eigenvectors of random matrices become uniformly distributed on the unit sphere as the dimension of the matrix is growing.

## 4. Increasing the number of generators

In [GJS] the following result was shown ([GJS, Prop. 2.1]):
Proposition 4.1. Let $\nu_{N, k}$ be the direct image of $d g_{1} d g_{2} \ldots d g_{k}$ on $G^{(k)}$ under the map

$$
\begin{equation*}
\left(g_{1}, g_{2}, \ldots, g_{k}\right) \rightarrow\left(\frac{1}{\sqrt{k}}\left(g_{1}+g_{1}^{-1}+\ldots+g_{k}+g_{k}^{-1}\right)\right)^{\wedge}\left(\pi_{N}\right) \tag{4.1}
\end{equation*}
$$

Thus $\nu_{N, k}$ is a probability measure on $\mathcal{H}_{N+1}$. As $k \rightarrow \infty, \nu_{N, k}$ converges in measure to the standard GOE measure on $\mathcal{H}_{N+1}$ if $N$ is even and to the standard GSE measure if $N$ is odd.

Here $\mathcal{H}_{N+1}$ denotes the real linear space of $(N+1) \times(N+1)$ real symmetric matrices.

Convergence of matrices implies convergence of their eigenvectors. Combining with results from [TV], we obtain the following results:

Theorem 4.2. Fix $N$ and let $k \rightarrow \infty$. Then the conclusions of [TV, Thm. 3 i)] and [TV, Corollary 4] holds for eigenvectors of $\left(g_{1}+g_{1}^{-1}+\ldots+g_{k}+\right.$ $\left.\left.g_{k}^{-1}\right)\right)^{\wedge}\left(\pi_{N}\right)$.

In other words, this implies that as $k \rightarrow \infty$, after a respective normalization of the eigenvectors (discussed in [TV]), the coefficients of the eigenvectors (for $M=o(\sqrt{N})$ being a coefficient of the eigenvectors) of $\pi_{N}(g)$ (after scaling them by $\sqrt{N}$ ), will differ from an independent random variable by $o(1)$ in the variation norm. This result is further studied in the numerical results shown later in the paper. A weaker result holds if $M=o(N / \log N)$.

## 5. Numerical implementation

In this section, we fix the number $k$ of generators of a subgroup of $\mathrm{SU}(2)$, and study numerically the distribution of the eigenvectors of 4.1.

The Lie algebra

$$
\operatorname{su}(2)=\left\{\left[\begin{array}{cc}
i a & z \\
-\bar{z} & -i a
\end{array}\right]: a \in \mathbb{R}, z \in \mathbb{C}\right\}
$$

is spanned by

$$
X_{1}=\left[\begin{array}{cc}
i & 0 \\
0 & -i
\end{array}\right], \quad X_{2}=\left[\begin{array}{cc}
0 & 1 \\
-1 & 1
\end{array}\right], \quad X_{3}=\left[\begin{array}{cc}
0 & i \\
i & 0
\end{array}\right] .
$$

Together with the exponential map and the differential of $\pi_{N}$, for $X \in \operatorname{su}(2)$, the relation

$$
\pi_{N}(\exp X)=\exp \left(d \pi_{N}(X)\right)
$$

is used to compute $\pi_{N}$ of some element of $\mathrm{SU}(2)$. First, we compute explicitely $d \pi_{N}$ of $X_{1}, X_{2}$ and $X_{3}$ with the relation $d \pi_{N}(X)=\left.\frac{d}{d t} \exp (t X)\right|_{t=0}$. This allows us to find $d \pi_{N}$ of any element of $\operatorname{su}(2)$. Then, the matrix exponential of Mathematica or Matlab is used to the matrix in $W_{N+1}$. The remaining of the section are the computations of $d \pi_{N}$ for $X_{1}, X_{2}, X_{3}$.

Starting with $X_{1}$, we have $\exp t X_{1}=\left[\begin{array}{cc}e^{t i} & 0 \\ 0 & e^{-t i}\end{array}\right]$. Since the action of
 of $W_{N+1}$, one gets, for $0<j<N$,

$$
\begin{aligned}
\left.\frac{d}{d t} \exp t X_{1}\left(e_{j}\right)\right|_{t=0} & =\left.\frac{d}{d t}\left(e^{i t} x\right)^{j}\left(e^{-i t} y\right)^{N-j}\right|_{t=0} \\
& =\left.\frac{d}{d t} e^{2 i t j-i t N} x^{j} y^{N-j}\right|_{t=0} \\
& =(2 i j-i N) e_{j}
\end{aligned}
$$

and $-i N e_{j}$ for $j=0, i N e_{j}$ for $j=N$.
Now, we have the computation of $\exp \left(t X_{2}\right)$. We notice that $\left(I_{2}\right.$ is the $2 \times 2$ identity matrix):

$$
\begin{gathered}
X_{2}^{2}=\left[\begin{array}{cc}
0 & 1 \\
-1 & 0
\end{array}\right]=-I_{2} \quad X_{2}^{3}=\left[\begin{array}{cc}
0 & -1 \\
1 & 0
\end{array}\right]=-X_{2} \\
X_{2}^{4}=\left[\begin{array}{ll}
1 & 0 \\
0 & 1
\end{array}\right]=I_{2} \quad X_{2}^{5}=\left[\begin{array}{cc}
0 & 1 \\
-1 & 0
\end{array}\right]=X_{2}
\end{gathered}
$$

This enables us to find:

$$
\begin{gathered}
\exp \left(t X_{2}\right)=I_{2}+t X_{2}-t^{2} \frac{I_{2}}{2!}-t^{3} \frac{X_{2}}{3!}+t^{4} \frac{I_{2}}{4!}+t^{5} \frac{X_{5}}{5!}+\ldots \\
\exp \left(t X_{2}\right)=\left[\begin{array}{cc}
\cos (t) & 0 \\
0 & \cos (t)
\end{array}\right]+\left[\begin{array}{cc}
0 & \sin (t) \\
-\sin (t) & 0
\end{array}\right]=\left[\begin{array}{cc}
\cos (t) & \sin (t) \\
-\sin (t) & \cos (t)
\end{array}\right]
\end{gathered}
$$

This is found by collecting terms and recognizing the infinite sums give the power series representation of $\sin (t)$ and $\cos (t)$.

Now, we have the following basis of $W_{n+1}$ (for $j=0 \ldots, N$ ):

$$
\hat{e}_{j}=\frac{x^{j} y^{N-j}}{\sqrt{j!(N-j)!}}
$$

We will now compute $d \pi_{N}\left(X_{2}\right)\left(\hat{e_{j}}\right)$. We claim that:

$$
d \pi_{N}\left(X_{2}\right)\left(\hat{e}_{j}\right)=-\sqrt{j(N-j+1)} \hat{e}_{j-1}+\sqrt{(N-j)(j+1)} \hat{e}_{j+1}
$$

Firstly, we have that the elements of the basis $\hat{e}_{j}$ are mapped to:

$$
\hat{e}_{j}=\frac{x^{j} y^{N-j}}{\sqrt{j!(N-j)!}} \rightarrow \frac{(x \cos (t)-y \sin (t))^{j}(x \sin (t)+y \cos (t))^{N-j}}{\sqrt{j!(N-j)!}}
$$

Now, it follows:

$$
\begin{equation*}
d \pi_{N}\left(X_{2}\right)\left(\hat{e}_{j}\right)=\left.\frac{d}{d t}\left(\frac{(x \cos (t)-y \sin (t))^{j}(x \sin (t)+y \cos (t))^{N-j}}{\sqrt{j!(N-j)!}}\right)\right|_{t=0} \tag{5.1}
\end{equation*}
$$

ĐANIEL CIZMA, JONATHAN GODIN, DMITRY JAKOBSON, AND LUIS LEDESMA VEGA
Assume now that $j \neq 0$ or $j \neq N$. Then, differentiating the above expression gives:

$$
\begin{aligned}
d \pi_{N}\left(X_{2}\right)\left(\hat{e}_{j}\right)= & \frac{1}{\sqrt{j!(N-j)!}}\left(\left(-j(x \cos (t)-y \sin (t))^{j-1}(x \sin (t)+y \cos (t))^{N-j+1}\right.\right. \\
& \left.\quad+(N-j)(x \sin (t)+y \cos (t))^{N-j-1}(x \cos (t)-y \sin (t))^{j+1}\right)\left.\right|_{t=0}
\end{aligned}
$$

Evaluating the following expression, this reduces to:

$$
d \pi_{N}\left(X_{2}\right)\left(\hat{e}_{j}\right)=\frac{-j}{\sqrt{j!(N-j)!}} x^{j-1} y^{N-(j-1)}+\frac{N-j}{\sqrt{j!(N-j)!}} y^{N-(j+1)} x^{j+1}
$$

By rearranging the terms, the result follows for $j \neq 0$ or $j \neq N$ :

$$
d \pi_{N}\left(X_{2}\right)\left(\hat{e}_{j}\right)=-\sqrt{j(N-j+1)} \hat{e}_{j-1}+\sqrt{(N-j)(j+1)} \hat{e}_{j+1}
$$

For $j=0, N$, the computations are very similar, and we find $\sqrt{N} \widehat{e}_{1}$ and $-\sqrt{N} \widehat{e}_{N-1}$ respectively.

To now compute $d \pi_{N}\left(X_{2}\right)$, simply notice that the action of $d \pi_{N}\left(X_{2}\right)$ on the basis vectors $\hat{e}_{j}$ will give us the columns in $d \pi_{N}\left(X_{2}\right)$. Hence, it follows:

$$
d \pi_{N}\left(X_{2}\right)\left(\hat{e}_{j}\right)= \begin{cases}\sqrt{N} \hat{e}_{1} & \text { if } j=0 \\ -\sqrt{j(N-j+1)} \hat{e}_{j-1}+\sqrt{(N-j)(j+1)} \hat{e}_{j+1} & \text { if } j \neq 0, j \neq N \\ -\sqrt{N} \hat{e}_{N-1} & \text { if } j=N\end{cases}
$$

For $X_{3}$, we note that $\left[\begin{array}{ll}0 & 1 \\ 1 & 0\end{array}\right]^{2}=\left[\begin{array}{ll}1 & 0 \\ 0 & 1\end{array}\right]=I_{2}$. Next, we have

$$
\begin{aligned}
\exp t X_{3} & =\sum_{n=0}^{\infty} \frac{t^{n} X_{3}^{n}}{n!} \\
& =\sum_{n=0}^{\infty} \frac{t^{n} i^{n}}{n!}\left[\begin{array}{ll}
0 & 1 \\
1 & 0
\end{array}\right]^{n} \\
& =I \sum_{n=0}^{\infty} \frac{(-1)^{2 n} t^{2 n}}{(2 n)!}+i\left[\begin{array}{ll}
0 & 1 \\
1 & 0
\end{array}\right] \sum_{n=0}^{\infty} \frac{(-1)^{2 n+1} t^{2 n+1}}{(2 n+1)!} \\
& =\left[\begin{array}{cc}
\cos t & i \sin t \\
i \sin t & \cos t
\end{array}\right]
\end{aligned}
$$

We know the acion of $\exp t X_{3}\left(\begin{array}{ll}x & y\end{array}\right) \mapsto(x \cos t+i y \sin t \quad i x \sin t+y \cos t)$. For $e_{j}$ with $1<j<N$, the computation becomes

$$
\begin{aligned}
& \left.\frac{d}{d t}\left(\exp t X_{3}\right)\left(e_{j}\right)\right|_{t=0} \\
& =\left.\frac{d}{d t}(x \cos t+i y \sin t)^{j}(i x \sin t+y \cos t)^{N-j}\right|_{t=0} \\
& =j(x \cos t+i y \sin t)^{j-1}(-x \sin t+i y \cos t)(i x \sin t+y \cos t)^{N-j} \\
& \quad \quad+\left.(x \cos t+i y \sin t)^{j}(N-j)(i x \sin t+y \cos t)^{N-j-1}(i x \cos t-y \cos t)\right|_{t=0}
\end{aligned}
$$

since $-x \sin t+i y \cos t=i(i x \sin t+y \cos t)$ and $i \cos t-y \sin t=i(x \cos t+i y \sin t)$,

$$
\begin{aligned}
& =\quad i j(x \cos t+i y \sin t)^{j-1}(i x \sin t+y \cos t)^{N-j+1} \\
& \quad \quad+\left.i(N-j)(x \cos t+i y \sin t)^{j+1}(i x \sin t+y \cos t)^{N-j-1}\right|_{t=0} \\
& =i j x^{j-1} y^{N-j+1}+i(N-j) x^{j+1} y^{N-j-1} .
\end{aligned}
$$

In the normalized base, we get

$$
\begin{aligned}
\left.\frac{d}{d t}\left(\exp t X_{3}\right)\left(\widehat{e}_{j}\right)\right|_{t=0}= & \left.\frac{1}{\sqrt{j!(N-j)!}} \frac{d}{d t}\left(\exp t X_{3}\right)\left(e_{j}\right)\right|_{t=0} \\
= & \frac{1}{\sqrt{j!(N-j)!}} i j x^{j-1} y^{N-j+1}+\frac{1}{\sqrt{j!(N-j)!}} i(N-j) x^{j+1} y^{N-j-1} \\
= & \frac{\sqrt{(j-1)!(N-j+1)!}}{\sqrt{j!(N-j)!}} \frac{i j x^{j-1} y^{N-j+1}}{\sqrt{(j-1)!(N-j+1)!}} \\
& +\frac{\sqrt{(j+1)!(N-j-1)!}}{\sqrt{j!(N-j)!}} \frac{i(N-j) x^{j+1} y^{N-j-1}}{\sqrt{(j+1)!(N-j-1)!}} \\
= & \sqrt{\frac{(N-j+1)}{j}} i j \widehat{e}_{j-1}+\sqrt{\frac{j+1}{(N-j)}} i(N-j) \widehat{e}_{j-1} \\
= & i \sqrt{j(N-j+1)} \widehat{e}_{j-1}+i \sqrt{(j+1)(N-j)} \widehat{e}_{j+1} .
\end{aligned}
$$

Lastly, for $j=0$ and $j=N$, one can find $d \pi_{N}\left(X_{3}\right)\left(\widehat{e}_{0}\right)=i \sqrt{N} \widehat{e}_{1}$ and $d \pi_{N}\left(X_{3}\right)\left(\widehat{e}_{N}\right)=i \sqrt{N} \widehat{e}_{N-1}$, respectively.

We conclude that

$$
d \pi_{N}\left(X_{3}\right)\left(\widehat{e}_{j}\right)= \begin{cases}i \sqrt{N} \widehat{e}_{1}, & \text { if } j=0 \\ i \sqrt{j(N-j+1)} \widehat{e}_{j-1}+i \sqrt{(j+1)(N-j)} \widehat{e}_{j+1}, & \text { if } 0<j<N \\ i \sqrt{N} \widehat{e}_{N-1}, & \text { if } j=N\end{cases}
$$

## 6. Testing uniform distribution: Even $N$

There are several ways of testing whether a sequence of points becomes uniformly distributed on $S^{N}$.

One can compute the $l^{k}, k \geq 2$ norms of eigenvectors, as well as their $l^{1}$ norm, and study its behaviour.

One can project eigenvectors onto an $m$-dimensional subspace of $\mathbb{R}^{N+1}$, and compare the distribution of the projections to the $m$-dimensional Gaussian.

Now, for $N$ even, the eigenvalues will be simple. Since $\pi_{N}(g)$ is selfadjoint and as the eigenvalues are simple, it follows that the eigenvectors will be real. Hence, we will plot the increasing $l^{p}$ norm of the eigenvectors of $\pi_{N}(g)$ for even $p$ against the eigenvalue number (in absolute value), and the same procedure will be done for the $l^{\infty}$ norm.


Figure 1. $l^{6}$ norm of the eigenvectors of $\pi_{N}(g)$ for $N=500$ against eigenvalue number (in absolute value)


Figure 2. $l^{8}$ norm of the eigenvectors of $\pi_{N}(g)$ for $N=500$ against eigenvalue number (in absolute value)


Figure 3. $l^{\infty}$ norm of the eigenvectors of $\pi_{N}(g)$ for $N=500$ against eigenvalue number (in absolute value)

The above graphs are to be expected, since $l^{p}$ is a decreasing function to $l^{\infty}$ as $p \rightarrow \infty$.


Figure 4. $l^{6}$ norm of the eigenvectors of $\pi_{N}(g)$ for $N=1000$ against eigenvalue number (in absolute value)


Figure 5. $l^{8}$ norm of the eigenvectors of $\pi_{N}(g)$ for $N=1000$ against eigenvalue number (in absolute value)


Figure 6. $l^{\infty}$ norm of the eigenvectors of $\pi_{N}(g)$ for $N=$ 1000 against eigenvalue number (in absolute value)


Figure 7. $l^{6}$ norm of the eigenvectors of $\pi_{N}(g)$ for $N=1500$ against eigenvalue number (in absolute value)


Figure 8. $l^{8}$ norm of the eigenvectors of $\pi_{N}(g)$ for $N=1500$ against eigenvalue number (in absolute value)


Figure 9. $l^{\infty}$ norm of the eigenvectors of $\pi_{N}(g)$ for $N=$ 1500 against eigenvalue number (in absolute value)

Now, we will make a histogram of the components of the eigenvectors of $\pi_{N}(g)$ (which are complex) and compare it against a histogram of a random variable distributed according to a bivariate normal distribution. The histogram of the eigenvectors will be made by taking all the eigenvectors and all of their components (making a 2 -dimensional histogram of the real and imaginary parts of the components of the eigenvectors). For the other histogram, we will take $(N+1)^{2}$ samples from a bivariate normal distribution.

Numerically, it was found that the components of the eigenvectors considered as a bivariate normal random variable (as a bivariate ( $X, Y$ ) random variable, where $X$ is the real part of the components of the eigenvectors and $Y$ the imaginary part) have an identity covariance matrix, so the histogram of the eigenvectors was compared against a bivariate normal distribution with 0 mean and an identity covariance matrix. To eliminate the bias present in the eigenvectors of the matrix, if $u=\left(u_{1}, \ldots, u_{N+1}\right)$ is a complex
eigenvector, we take $\arg u_{1}$ and multiply the eigenvector by $e^{-i \arg \left(u_{1}\right)}$. We will multiply the components by $\sqrt{2(N+1)}$ (in order to obtain a normal distribution). In both figures the generators were fixed and $N$ increased.


Figure 10. Histogram of the bivariate gaussian distribution against the histogram of the eigenvector components of $\pi_{N}(g)$ for $N=600$


Figure 11. Histogram of the bivariate gaussian distribution against the histogram of the eigenvector components of $\pi_{N}(g)$ for $N=1000$

## 7. Testing uniform distribution: odd $N$

For odd $N$, the eigenvalues will be double ( $\pi_{N}(g)$ has even dimension). In a similar way to the histograms produced for the even case, we do the same normalization process for the eigenvectors. Now, for each eigenvalue $\lambda_{i}$, we will have two associated eigenvectors (which form a basis for the eigenspace associated to $\lambda_{i}$ ) $u_{i}$ and $u_{i+1}$. Then, consider the random variable $(X, Y, Z, W)$, where we have:

$$
X=\operatorname{Re}\left(u_{i_{j}}\right) \quad Y=\operatorname{Im}\left(u_{i_{j}}\right) \quad Z=\operatorname{Re}\left(u_{i+1_{j}}\right) \quad X=\operatorname{Im}\left(u_{i+1_{j}}\right)
$$

For $u_{i_{j}}$ being the $j$-th component of $u_{i}$. Numerically, the random variable ( $X, Y, Z, W$ ) seems to have 0 mean and an identity covariance matrix as $N$ becomes large. We consider the random variables coming from projections of $(X, Y, Z, W),(X, Y),(X, W),(Z, Y),(Z, W)$ and fixed generators for $\pi_{N}(g)$. The histogram for the projected random variables will be compared against a histogram of $\frac{(N+1)^{2}}{4}$ samples from a bivariate standard normal distribution (the projection random variables had as well 0 mean and an identity covariance matrix).

(a) Histogram of the bivariate gaussian for 601

(b) Histogram of the bivariate gaussian for $N=1001$

Figure 12. Histogram of the bivariate standard gaussian distribution for samples corresponding to $N=601$ and $N=$ 1001


Figure 13. Histograms of projections of $(X, Y, Z, W)$ for $N=601$


Figure 14. Histograms of projections of $(X, Y, Z, W)$ for $N=1001$

## 8. Ramanujan Bounds

Another interesting point is to analyze the limiting behavior of the eigenvectors of $\pi_{N}(g)$ for different number of generators (which are randomly chosen). It is expected that the largest eigenvalues (in absolute value) of $\pi_{N}(g)$ for any number $k$ of generators tend to $\pm 2 \sqrt{2 k-1}$ as $n \rightarrow \infty$, also known as the Ramanujan bound. This assertion has an equivalent for random graphs, also known as Alon's conjucture, proven recently.

The statement will be verified numerically for $N$ in a range between 699 and 799 , for a varying number of generators (specifically, for $2,3,4$ and 5 generators). The largest eigenvalue number will be compared against $N$, where the straight lines in the graphs represent the Ramanujan bound.


Figure 15. Largest eigenvalue number of $\pi_{N}(g)$ for $k=2,3,4,5$ generators chosen at random

## 9. Representation varieties

Moving in the space of representations. LATER ON.

## 10. The Spectral Function

It is also interesting to look at the behaviour of the spectral function

$$
\begin{equation*}
F(i, j, \mu)=\sum_{\lambda_{m} \leq \mu} u_{m}(i) \overline{u_{m}(j)}, \tag{10.1}
\end{equation*}
$$

where $\lambda_{m}, 1 \leq m \leq N+1$, are the eigenvalues of $\pi_{N}(g)$ and $u_{m}(i), u_{m}(j)$ are the $i$-th and $j$-th component, respectively, of the corresponding eigenvector. Note that the spectral function defined in 10.1) is "self-normalizing" (as a kernel for a spectral projection operator), in the sense of Definition (i) from TV].


Figure 17. Shown above are plots of the spectral function, $F(i, j, \mu)$ of $\pi_{N}(g), N=2000, g \in \mathbb{R}[\mathrm{SU}(2)]$ fixed, for various $i$ and $j$ : (a) $F(18,18, \mu)$ (b) $F(0,1, \mu)$ (c) $F(13,23, \mu)(\mathrm{d})$ $F(13,113, \mu)$

Thing to do: look at matrix entries of $\pi_{N}(g)$, and see if one can notice anything interesting.

One way to gain qualitative information about $\pi_{N}(g)$ is to make a color map of the matrix. Considering a real matrix $A_{N}(g)$, whose entries are related to the entries of $\pi_{N}(g)$ by some scalar function (modulus, real part etc.), we can color an $N \times N$ grid according to the entries of $A_{N}(g)$ and some color scale.

$$
B=\left(\begin{array}{lllll}
1 & 0 & 0 & 0 & 0 \\
0 & 2 & 0 & 0 & 0 \\
0 & 0 & 3 & 0 & 0 \\
0 & 0 & 0 & 4 & 0 \\
0 & 0 & 0 & 0 & 5
\end{array}\right)
$$



Figure 18. Example of a color map of the matrix $B$ shown above.

We can apply this to a matrix $A_{N}(g)$, where we take the entries of $A_{N}(g)$ to be the modulus of the entries of $\pi_{N}(g)$. Doing this gives some interesting results.


Figure 19. Color maps of the matrix $A_{N}(g)$ for various $g \in \mathbb{R}[\mathrm{SU}(2)]$ and $N \in \mathbb{N}$.

While different $g \in \mathbb{R}[\mathrm{SU}(2)]$ produce different maps, they all seem to share the similarity of having a family of ellipses inscribed within the matrix grid.

## References

[GJS] A. Gamburd, D. Jakobson and P. Sarnak. Spectra of elements in the group ring of SU(2). J. Eur. Math. Soc. 1 (1999), 51-85.
[KnY] A. Knowles and J. Yin. Eigenvector distribution of Wigner matrices. Probab. Theory Related Fields 155 (2013), no. 3-4, 543-582.
[TV] T. Tao and V. Vu. Random matrices: universal properties of eigenvectors. Random Matrices Theory Appl. 1 (2012), no. 1, 1150001.

E-mail address: daniel.cizma@mail.mcgill.ca
E-mail address: jonathan.godin@mail.mcgill.ca
E-mail address: jakobson@math.mcgill.ca
E-mail address: luis.ledesmavega@mail.mcgill.ca
Dept. of Mathematics and Statistics, McGill university, 805 Sherbrooke West, Montreal, Quebec, H3A0B9, Canada

